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On level crossings
of certain Markovian jump processes

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1. The process

We consider a real-valued piecewise-deterministic Markov process $X = (X_t)_{t \geq 0}$ driven by three characteristics:

- (i) a jump intensity function $\lambda : \mathbb{R} \rightarrow [0, \infty)$,
- (ii) a drift coefficient $\mu : \mathbb{R} \rightarrow \mathbb{R}$,
- (iii) and a stochastic kernel $J(x, dz)$ from \mathbb{R} to \mathbb{R} satisfying $J(x, \{0\}) = 0$.

Interpretation of the characteristics:

- (i) The function λ gives the stochastic intensity of jumps:

$$\mathbb{P}(X \text{ jumps in } [t, t+h] | \mathcal{F}_t^X) = \lambda(X_t)h + o(h).$$

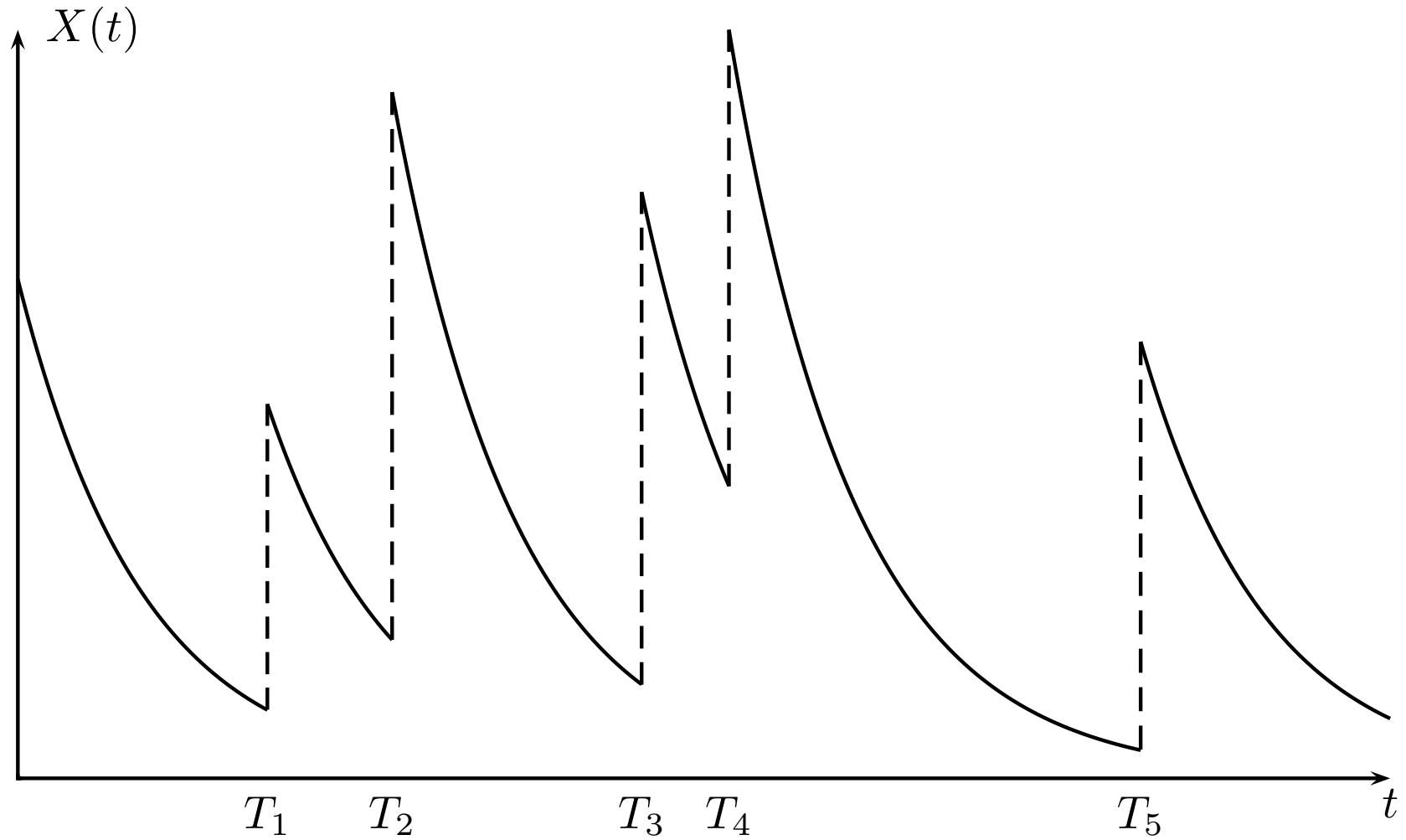
- (ii) The function μ describes the deterministic evolution of X between the jumps:

$$\frac{d}{dt}X_t = \mu(X_t).$$

- (iii) The kernel J is the conditional jump size distribution:

$$\mathbb{P}(X_t - X_{t-} \in \cdot | \mathcal{F}_{t-}^X, X \text{ jumps at } t) = J(X_{t-}, \cdot).$$

A possible trajectory in case $\mu(x) = -x$



Assumptions:

- (i) The function μ is right-continuous and the set

$$D_\mu := \{x \in \mathbb{R} : \mu(x) = 0\}$$

is locally finite. For any $x \in \mathbb{R}$, there exists a unique continuous function $q(x, \cdot) : [0, \infty) \rightarrow \mathbb{R}$ satisfying the integral equation

$$q(x, t) = x + \int_0^t \mu(q(x, s)) ds, \quad t \geq 0.$$

- (ii) The jump intensity λ is assumed to be measurable, locally bounded and such that

$$\int_0^\infty \lambda(q(x, s)) ds = \infty, \quad x \in \mathbb{R}.$$

Applications:

- storage processes (Harrison and Resnick 1976, Meyn and Tweedie 1993)
- stress release models (Vere-Jones 1988, Zheng 1991, Borovkov and Vere-Jones 2000, Last 2004)
- queueing models (Browne and Sigman 1992, Meyn and Tweedie 1993)
- repairable systems (Last and Szekli 1998)

Definition: (X_t) is the homogeneous Markov process w.r.t. a family $\{\mathbb{P}_x : x \in \mathbb{R}\}$ of probability measures on (Ω, \mathcal{F}) satisfying

$$\begin{aligned} \mathbb{E}_x f(X_{t \wedge \tau_m}) &= f(x) + \mathbb{E}_x \int_0^{t \wedge \tau_m} f'(X_s) \mu(X_s) ds \\ &\quad + \mathbb{E}_x \int_0^{t \wedge \tau_m} \int_{\mathbb{R}} (f(X_s + z) - f(X_s)) \lambda(X_s) J(X_s, dz) ds. \end{aligned}$$

for all $x \in \mathbb{R}$, $m \in \mathbb{N}$, and all absolutely continuous $f : \mathbb{R} \rightarrow \mathbb{R}$ such that the Radon–Nikodym derivative f' as well as the function

$$x \mapsto \lambda(x) \int (f(x + z) - f(x)) J(x, dz)$$

be locally bounded. Here,

$$\tau_m := \inf\{t \geq 0 : |X(t)| \geq m\}, \quad m \in \mathbb{N}.$$

Remark: The infinitesimal generator \mathcal{A}_m of $(X_{t \wedge \tau_m})$ is given by

$$\mathcal{A}_m f(x) = f'(x)\mu(x) + \lambda(x) \int_{\mathbb{R}} (f(x+z) - f(x))J(x, dz), \quad |x| < m.$$

Remark: Let $N := (T_n)_{n \geq 1}$ be the point process of jump times of X and let (\mathcal{F}_t) be the filtration generated by X . Then N has stochastic intensity $\lambda(X_{t-})$. More generally, the marked point process

$$\Phi := ((T_n, X_{T_n} - X_{T_n-}))_{n \geq 1}$$

has the compensator

$$\nu(d(t, z)) := \lambda(X_t)J(X_t, dz)dt$$

Assumptions:

- (i) We have $\mathbb{P}_x(\lim_{n \rightarrow \infty} T_n = \infty) = 1$ for all $x \in \mathbb{R}$.
- (ii) The process (X_t) has an invariant distribution π satisfying

$$\lambda_\pi := \mathbb{E}_\pi N(1) < \infty,$$

where $N(1)$ is the number of jumps in the time interval $[0, 1]$.

Remark: Assume that there is an $\varepsilon > 0$ such that

$$\liminf_{x \rightarrow -\infty} (\mu(x) + \lambda(x)m^+(x)(1 - \varepsilon) - \lambda(x)m^-(x)) > 0,$$

$$\limsup_{x \rightarrow \infty} (\mu(x) + \lambda(x)m^+(x) - \lambda(x)m^-(x)(1 - \varepsilon)) < 0,$$

where

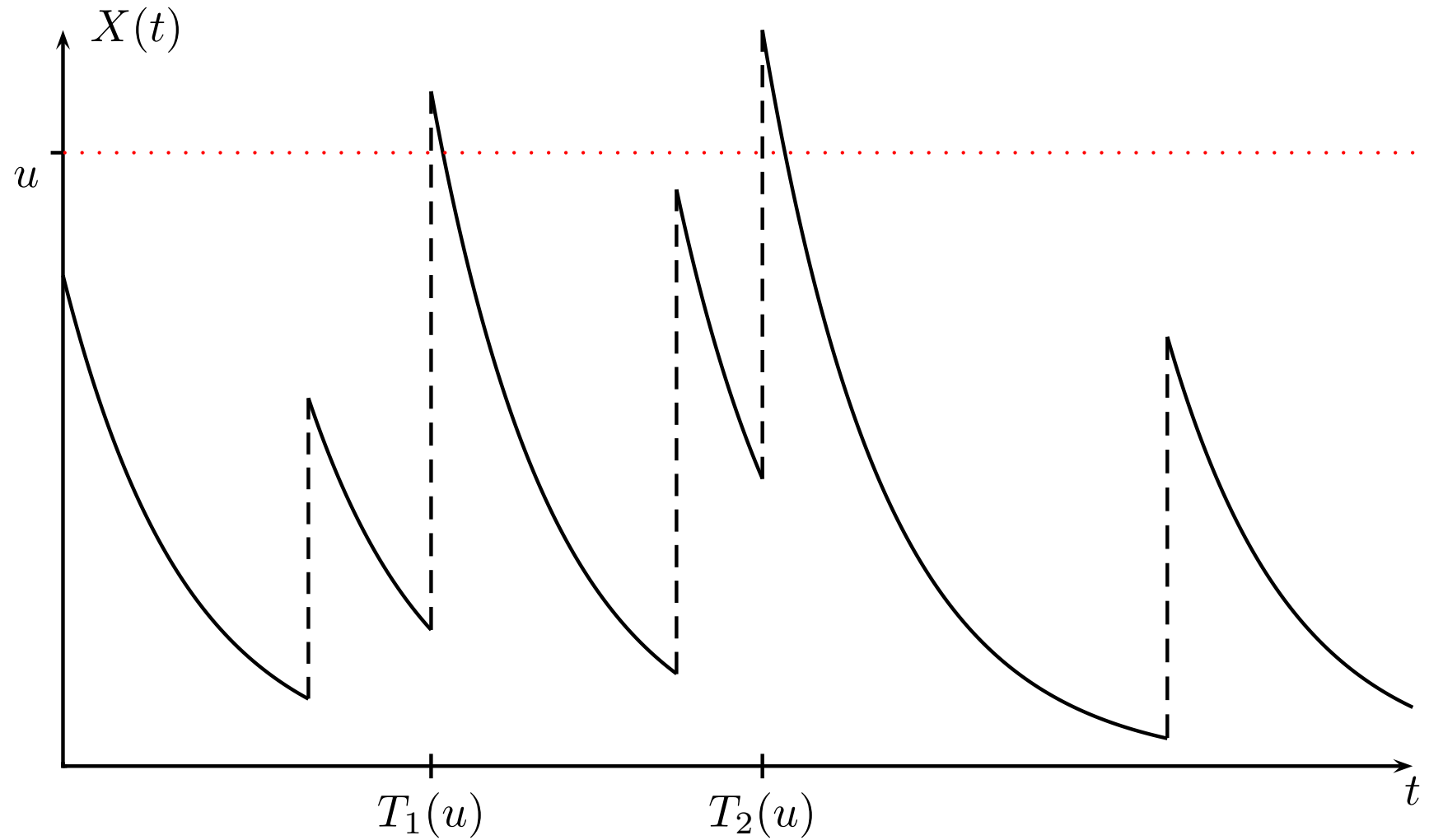
$$m^-(x) := - \int_{-\infty}^0 zJ(x, dz), \quad m^+(x) := \int_0^{\infty} zJ(x, dz), \quad x \in \mathbb{R}.$$

Under further technical assumptions there is a unique invariant distribution π . The condition $\lambda_\pi < \infty$ requires a slight strengthening of the above ergodicity assumption.

2. Up- and downcrossings

Definition: The process (X_t) has an *upcrossing* of level $u \in \mathbb{R}$ at time $s > 0$ if there is some $\delta > 0$ such that $X_t < u$ for $s - \delta \leq t < s$ and $X_t \geq u$ for $s < t \leq s + \delta$. If, in addition, $X_{s-} = X_s (= u)$ then the upcrossing is *continuous*. Otherwise it is *discontinuous*. *Downcrossings* are defined similarly.

Definition: The point process of all times of continuous up- and downcrossings of level u is denoted by N^u . The point process of all times of upcrossings of level u is denoted by N_+^u .

Upcrossings of level u 

Remark: Under the stationary probability measure

$$\mathbb{P} := \int \mathbb{P}_x \pi(dx)$$

the point processes N^u and N_+^u are (jointly) stationary. Their intensities are denoted by $\nu(u)$ and $\nu_+(u)$ respectively.

Theorem: (Rice's formula) *The stationary distribution π is absolutely continuous on $\mathbb{R} \setminus D_\mu$. The density p satisfies*

$$\nu(u) = |\mu(u)|p(u), \quad u \notin D_\mu.$$

3. Compound Poisson limit

Definition: Let $\rho \in [0, 1)$. A geometrically compound Poisson process is a point process with independent increments defined as follows. Each point of a homogeneous Poisson process of intensity $(1 - \rho)$ gets (independently of the other points) a mass $k \in \{1, 2, \dots\}$ with probability $(1 - \rho)\rho^{k-1}$.

Remark: As the above geometric distribution has mean $1/(1 - \rho)$, the intensity of Π_ρ is 1.

Remark: Compound Poisson limits occur quite frequently in extreme value theory, see e.g. Leadbetter and Hsing (1990) or the book Falk, Hüsler and Reiss (2004).

Assumption: We have $\nu_+(u) > 0$ for all sufficiently large $u \in \mathbb{R}$.

Theorem: Consider, for $b \in \mathbb{R}$, the time-scaled level crossing process

$$M^b(t) := N_+^b(\nu_+(b)^{-1}t), \quad t \geq 0.$$

Under the scenarios described below we have weak convergence

$$\mathbb{P}(M^b \in \cdot) \xrightarrow{w} \mathbb{P}(\Pi_\rho \in \cdot) \quad \text{as } b \rightarrow \infty.$$

The number $\rho \in [0, 1)$ is determined by the characteristics of (X_t) .

More generally, we have weak convergence of $\mathbb{P}_x(M^b \in \cdot)$ for any $x \in \mathbb{R}$ that is attracted by the support of π .

Assumptions: We assume that one of the following assumptions is satisfied.

- (i) We have $J(x, (0, \infty)) = 1$ and $\mu(x) < 0$ for all $x > 0$.
- (ii) We have $J(x, (-\infty, 0)) = 1$ for all $x \in \mathbb{R}$ and $\mu(x) > 0$ for all $x > 0$.

Remark: It is sufficient to make this assumptions only in an asymptotic sense.

Definition: Let

$$\tau(u) := \inf\{t > 0 : N^u(t) \geq 1\}, \quad u \in \mathbb{R},$$

be the time of the first continuous crossing of level u . Define

$$\gamma(u, b) := \mathbb{P}_b(\tau(b) < \tau(u)), \quad u < b,$$

as the probability that the process, when starting at level b (continuously) crosses level b before dropping to level u .

Theorem: *Assume that*

$$\lim_{u \rightarrow \infty} \liminf_{b \rightarrow \infty} \gamma(u, b) = \lim_{u \rightarrow \infty} \limsup_{b \rightarrow \infty} \gamma(u, b) = \rho$$

for some $\rho \in [0, 1)$. Then the compound Poisson limit theorem holds.

4. Scenarios for the compound Poisson limit

Assumption (A1):

- (i) $\mu(y) \rightarrow \mu(\infty) < 0$ and $\lambda(y) \rightarrow \lambda(\infty) \in [0, \infty)$ as $y \rightarrow \infty$.
- (ii) There are random variables $\bar{\xi}(u)$ and $\underline{\xi}(u)$, $u > 0$, such that $\mathbb{E}\bar{\xi}(u) < \infty$ and

$$\underline{\xi}(u) \stackrel{d}{\leq} \xi(x) \stackrel{d}{\leq} \bar{\xi}(u), \quad x \geq u > 0,$$

where $\xi(x)$ has distribution $J(x, \cdot)$.

- (iii) As $y \rightarrow \infty$ we have that $\bar{\xi}(y), \underline{\xi}(y) \xrightarrow{d} \xi(\infty)$, where $\xi(\infty)$ is an integrable random variable satisfying

$$\mathbb{E}\xi(\infty) + \mu(\infty)/\lambda(\infty) < 0.$$

Lemma: *If assumption (A1) holds then*

$$\lim_{u \rightarrow \infty} \liminf_{b \rightarrow \infty} \gamma(u, b) = \lim_{u \rightarrow \infty} \limsup_{b \rightarrow \infty} \gamma(u, b) = \rho,$$

where $\rho \in [0, 1)$ is given by

$$\rho = -\frac{\lambda(\infty)}{\mu(\infty)} \mathbb{E}\xi(\infty).$$

Remark: ρ is the probability that a stationary $M/GI/1/\infty$ -queue with arrival intensity $-\lambda(\infty)/\mu(\infty)$ and generic service time $\xi(\infty)$ is non-empty.

Assumption (A2):

- (i) $\mu(y) \rightarrow \mu(\infty) > 0$ and $\lambda(y) \rightarrow \lambda(\infty) \in [0, \infty)$ as $y \rightarrow \infty$.
- (ii) There are random variables $\bar{\xi}(u)$ and $\underline{\xi}(u)$, $u > 0$, such that $\mathbb{E}\bar{\xi}(u) < \infty$ and

$$\underline{\xi}(u) \stackrel{d}{\leq} \xi(x) \stackrel{d}{\leq} \bar{\xi}(u), \quad x \geq u > 0,$$

where $\xi(x)$ has distribution $J(x, \cdot)$.

- (iii) As $y \rightarrow \infty$ we have that $\bar{\xi}(y), \underline{\xi}(y) \xrightarrow{d} \xi(\infty)$, where $\xi(\infty)$ is an integrable random variable satisfying

$$\mathbb{E}\xi(\infty) + \mu(\infty)/\lambda(\infty) < 0.$$

Lemma: *If assumption (A2) holds then*

$$\lim_{u \rightarrow \infty} \liminf_{b \rightarrow \infty} \gamma(u, b) = \lim_{u \rightarrow \infty} \limsup_{b \rightarrow \infty} \gamma(u, b) = \rho,$$

where

$$\rho = 1 - \frac{w\mu(\infty)}{\lambda(\infty)}$$

and w is the unique solution of

$$\mathbb{E}e^{w\xi(\infty)} = 1 - w\mu(\infty)/\lambda(\infty).$$

Remark: ρ is the probability that a typical customer, arriving at a stationary $GI/M/1/\infty$ -queue with service intensity $\lambda(\infty)/\mu(\infty)$ and generic interarrival time $-\xi(\infty)$, finds the system non-empty.

Assumption (A3):

- (i) $\mu(y) \rightarrow -\infty$ as $y \rightarrow \infty$.
- (ii) There are random variables $\bar{\xi}(u)$, $u > 0$, such that $\mathbb{E}\bar{\xi}(u) < \infty$ and

$$\xi(x) \stackrel{d}{\leq} \bar{\xi}(u), \quad x \geq u > 0,$$

where $\xi(x)$ has distribution $J(x, \cdot)$.

- (iii) We have

$$\mathbb{E}\bar{\xi}(u_0) + \bar{\mu}(u_0)/\bar{\lambda}(u_0) < 0,$$

for some $u_0 > 0$, where

$$\bar{\mu}(u_0) := \sup_{x \geq u_0} \mu(x), \quad \bar{\lambda}(u_0) := \sup_{x \geq u_0} \lambda(x).$$

Assumption (A4):

- (i) $\lambda(y) \rightarrow \infty$ as $y \rightarrow \infty$.
- (ii) There are random variables $\underline{\xi}(u)$, $u > 0$, such that $\mathbb{E}\bar{\xi}(u) < \infty$ and

$$\xi(x) \stackrel{d}{\leq} \bar{\xi}(u), \quad x \geq u > 0,$$

where $\xi(x)$ has distribution $J(x, \cdot)$.

- (iii) We have

$$\mathbb{E}\bar{\xi}(u_0) + \bar{\mu}(u_0)/\underline{\lambda}(u_0) < 0,$$

for some $u_0 > 0$, where $\underline{\lambda}(u_0) := \inf_{x \geq u_0} \lambda(x)$.

Lemma: *If assumption (A3) or assumption (A4) holds, then*

$$\lim_{b \rightarrow \infty} \gamma(u, b) = 0$$

for all sufficiently large u .

Remark: If assumption (A3) or assumption (A4) holds, then (X_t) drops very quickly from a high level.

Remark: If one of the Assumptions (A1)–(A4) holds, then $\nu_+(b) = \nu(b)$ for all $b > 0$. By Rice's formula this means that

$$\nu_+(b) = |\mu(b)|p(b), \quad b > 0.$$

5. Asymptotics for the first crossing times

Theorem: *Let $b \in \mathbb{R}$ and define the first and second upward crossing time of level b*

$$T_1(b) := \inf\{t > 0 : X_t \geq b\},$$

$$T_2(b) := \inf\{t > T_1(b) : X_t \geq b, X_{t-\delta} < b \text{ for some } \delta > 0\}.$$

If one of the Assumptions (A1)–(A4) holds, then, as $b \rightarrow \infty$,

$$\mathbb{P}_x((1 - \rho)\nu_+(b)T_1(b) > s) \rightarrow e^{-s},$$

$$\mathbb{P}_x((1 - \rho)\nu_+(b)T_2(b) > s) \rightarrow e^{-s}(1 + (1 - \rho)s)$$

for any x that is attracted by π .