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Filtering of marked point processes

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Reference:

Last, G. and Brandt, A. (1995) *Marked Point Processes on the Real Line: The Dynamic Approach*. Springer-Verlag, New York.

1. Marked point processes

Let \mathbf{X} be a Polish space and x_∞ a point external to \mathbf{X} . A **marked point process** (MPP) with mark space \mathbf{X} is a sequence

$$\Phi = ((T_n, X_n))_{n \geq 1}$$

of random times $T_n \in (0, \infty]$ and random $(\mathbf{X} \cup \{x_\infty\})$ -valued marks X_n such that

$$\begin{aligned} T_n < T_{n+1} & \quad \text{if } T_n < \infty, \\ T_n = T_{n+1} = \infty & \quad \text{if } T_n = \infty, \end{aligned}$$

and

$$\{T_n = \infty\} = \{X_n = x_\infty\}.$$

A MPP Φ is identified with a random measure on $[0, \infty) \times \mathbf{X}$:

$$\Phi \equiv \sum_{n \geq 1} \mathbf{1}\{T_n < \infty\} \delta_{(T_n, X_n)}.$$

For any $t \geq 0$, Φ_t and Φ_{t-} denote the restrictions of Φ to $[0, t] \times \mathbf{X}$ resp. $[0, t) \times \mathbf{X}$. We denote

$$\mathcal{F}_t^\Phi := \sigma(\Phi_t), \quad \mathcal{F}_{t-}^\Phi := \sigma(\Phi_{t-}).$$

$\{\mathcal{F}_t^\Phi : t \geq 0\}$ is the **filtration** generated by Φ . Given any filtration $\{\mathcal{F}_t : t \geq 0\}$ satisfying

$$\mathcal{F}_t^\Phi \subset \mathcal{F}_t, \quad t \geq 0,$$

the $\{\mathcal{F}_t\}$ -**compensator** is an $\{\mathcal{F}_t\}$ -predictable random measure ν on $[0, \infty) \times \mathbf{X}$ satisfying the heuristic relationship

$$\nu(d(t, x)) = P(\Phi(d(t, x)) > 0 | \mathcal{F}_{t-}).$$

2. Detecting change points

Let $\tau \in (0, \infty]$ be a random time with c.d.f. F having a density f (on $(0, \infty)$) and let $\lambda_1, \lambda_2 > 0$ with $\lambda_1 < \lambda_2$. Assume that $\Phi = (T_n)$ is a **doubly stochastic Poisson process** with respect to τ whose intensity switches from λ_1 to λ_2 at time τ . Hence, given τ , Φ is a Poisson process with intensity

$$\lambda(t) := \mathbf{1}\{t < \tau\}\lambda_1 + \mathbf{1}\{t \geq \tau\}\lambda_2, \quad t \geq 0.$$

Filtering problem:

Assume that Φ can be observed up to time t .

- Has the intensity switched by time t ?
- When did the change happen?

Joint distribution of Φ_t and τ :

$$\mathbf{1}\{s \leq t\} \mathbb{P}(\Phi_t \in d\varphi, \tau \in ds) = \mathbf{1}\{s \leq t\} L(\varphi, s, t) f(s) e^t \mathbb{Q}^t(d\varphi) ds,$$

$$\mathbb{P}(\Phi_t \in d\varphi, \tau > t) = L(\varphi, t, t) (1 - F(t)) f(s) e^t \mathbb{Q}^t(d\varphi),$$

where

$$L(\varphi, s, t) := \lambda_1^{\varphi(s)} e^{-\lambda_1 s} \lambda_2^{\varphi(t) - \varphi(s)} e^{-\lambda_2(t-s)},$$

$\varphi(s)$ is the number of points of φ in $[0, s]$ and \mathbb{Q}^t is the distribution of a homogeneous Poisson process on $[0, t]$ of intensity 1.

Theorem: *We have*

$$\mathbf{1}\{s \leq t\} \mathbb{P}(\tau \in ds | \Phi_t) = \frac{L(\Phi, s, t) f(s) ds}{\int_0^t L(\Phi, u, t) f(u) du + L(\Phi, t, t)(1 - F(t))}$$

and in particular

$$\hat{Z}(t) := \mathbb{P}(\tau \leq t | \Phi_t) = \frac{\int_0^t L(\Phi, s, t) f(s) ds}{\int_0^t L(\Phi, u, t) f(u) du + L(\Phi, t, t)(1 - F(t))}.$$

Moreover, for $T_n < t < T_{n+1}$ we have

$$\hat{Z}(t) = \hat{Z}(T_n) + \int_{T_n}^t r(s)(1 - \hat{Z}(s)) ds + \int_{T_n}^t (\lambda_1 - \lambda_2) \hat{Z}(s)(1 - \hat{Z}(s)) ds,$$

where $r(s) := f(s)/(1 - F(s))$ is the hazard rate of F , and

$$\hat{Z}(T_n) = \frac{\lambda_2 \hat{Z}(T_n^-)}{\lambda_1 + (\lambda_2 - \lambda_1) \hat{Z}(T_n^-)}, \quad n \geq 1.$$

Stochastic $\{\mathcal{F}_t\}$ -intensity of Φ :

$$\begin{aligned}\hat{\lambda}(t) &:= \mathbb{E}[\lambda(t)|\Phi_{t-}] = (1 - \hat{Z}(t-))\lambda_1 + \hat{Z}(t-)\lambda_2 \\ &= \lambda_1 + (\lambda_2 - \lambda_1)\hat{Z}(t-).\end{aligned}$$

Semimartingale representation of $\hat{Z}(t)$:

$$\hat{Z}(t) = \int_0^t r(s)(1 - \hat{Z}(s)) ds + \int_0^t Y_s d\hat{M}(s)$$

where

$$Y_s = \frac{(\lambda_2 - \lambda_1)\hat{Z}(s-)(1 - \hat{Z}(s-))}{\lambda_1 + (\lambda_2 - \lambda_1)\hat{Z}(s-)}$$

and

$$\hat{M}(t) := \Phi(t) - \int_0^t \hat{\lambda}(s) ds$$

is the **innovations martingale**.

2. Filtering of Semi-Markov processes

Consider a MPP $\Psi = ((T'_n, Y_n))$ with finite mark space $\{1, \dots, m\}$. Let $T'_0 \equiv 0$ and Y_0 be a random element of $\{1, \dots, m\}$. Let

$$(T^t, Z(t)) := (T'_n, Y_n) \quad \text{if} \quad T'_n \leq t < T'_{n+1}.$$

Define

$$\mathcal{F}_t := \sigma(Y_0, \Psi_t), \quad t \geq 0,$$

and assume that the $\{\mathcal{F}_t\}$ -compensator ν of Ψ is given by

$$\nu(dt \times \{j\}) = r_{Z(t^-)j}(t - T^{t^-})dt,$$

where the r_{ij} are measurable functions such that

$$r_{ii}(t) \equiv 0, \quad i = 1, \dots, m.$$

We can assume that

$$Y_n \neq Y_{n+1} \quad \text{if} \quad T'_n < \infty.$$

Let

$$q_{ij}(t) := r_{ij}(t) \exp \left[- \int_0^t \sum_k r_{ik}(s) ds \right], \quad t \geq 0.$$

We allow for

$$\sum_{j=1}^m \int_0^\infty r_{ij}(s) ds < \infty,$$

i.e.

$$\sum_{j=1}^m \int_0^\infty q_{ij}(s) ds < 1,$$

and define

$$\bar{Q}_i(t) := 1 - \sum_{j=1}^m \int_0^t q_{ij}(s) ds, \quad t \geq 0.$$

We assume that $\Phi = (T_n)$ is a **doubly stochastic Poisson process** with respect to $\{Z(t)\}$. More precisely we consider given intensities

$$\lambda_1, \dots, \lambda_m > 0$$

and assume that the **conditional distribution** of Φ given the process $\{Z(t)\}$ is that of a Poisson process with intensity function

$$t \mapsto \lambda_{Z(t)}.$$

Let Z' be a $\sigma(Y_0)$ -measurable random element of some measurable space and define the **observations**

$$\mathcal{G}_t := \sigma(Z', \Phi_t), \quad t \geq 0.$$

Theorem: *We have*

$$\mathbb{P}(T^t \in ds, Z(t) = j | \mathcal{G}_t) = \frac{Q(t; j, ds)}{\sum_k Q(t; k, [0, \infty] \times \mathbf{Y})} \quad \mathbb{P} - a.s.,$$

where

$$\begin{aligned} Q(t; j, ds) := & \mathbb{P}(Y_0 = j | Z') \lambda_j^{\Phi(t)} e^{-\lambda_j t} \bar{Q}_j(t) \delta_0(ds) \\ & + \mathbf{1}\{s \leq t\} \lambda_j^{\Phi(t) - \Phi(s)} e^{-\lambda_j(t-s)} \bar{Q}_j(t-s) p(s, j) ds \end{aligned}$$

and $p : \Omega \times [0, \infty) \times \{1, \dots, m\} \rightarrow [0, \infty)$ is a $\{\mathcal{G}_t\}$ -optional process satisfying

$$\begin{aligned} p(t, j) = & \sum_{k=1}^m \lambda_k^{\Phi(t)} e^{-\lambda_k t} q_{kj}(t) \mathbb{P}(Y_0 = k | Z') \\ & + \sum_{k=1}^m \int_0^t \lambda_k^{\Phi(t) - \Phi(u)} e^{-\lambda_k(t-u)} q_{kj}(t-u) p(u, k) du. \end{aligned}$$

4. Generalizations

Consider a MPP $\Psi = ((T'_n, Y_n))$ with **discrete** mark space \mathbf{Y} and let Y_0 be a random element of \mathbf{Y} . Denote

$$\mathcal{F}_t := \sigma(Y_0, \Psi_t), \quad t \geq 0,$$

and define the $\{\mathcal{F}_t\}$ -adapted process $\{Y(t)\}$ by

$$Y(t) := Y_n \quad \text{if } T'_n \leq t < T'_{n+1}$$

and $Y(t) := y_\infty$ if $t \geq T'_\infty := \lim_{n \rightarrow \infty} T'_n$. (Here $T'_0 := 0$.) Let

$$g_1 : [0, \infty) \times \mathbf{Y} \times [0, \infty) \times \mathbf{Y} \rightarrow \{0, 1\}$$

be measurable and Ψ^0 the unique PP satisfying

$$\Psi^0(\cdot) = \int \mathbf{1}\{g_1(T^{t-}(\Psi^0), Y(t-), t, y) = 0\} \mathbf{1}\{t \in \cdot\} \Psi(d(t, y)),$$

where $T^{t-}(\Psi^0)$ is the last point of Ψ^0 strictly before t ($= 0$ if there is no such point).

The **observed** MPP $\Phi = ((T_n, X_n))$ with **discrete** mark space \mathbf{X} is assumed to be given by

$$\Phi = \int g_1(\pi^{t-}(\Psi^0), Y(t-), t, y) \mathbf{1}\{g_2(T^{t-}(\Psi^0), Y(t-), t, y) = x\} \mathbf{1}\{(t, x) \in \cdot\} \Psi(d(t, y)),$$

where

$$g_2 : [0, \infty) \times \mathbf{Y} \times [0, \infty) \times \mathbf{Y} \rightarrow \mathbf{X}$$

is a given measurable function. Let Z' be a $\sigma(Y_0)$ -measurable random element of some measurable space and define the **observations** by

$$\mathcal{G}_t := \sigma(Z', \Phi_t), \quad t \geq 0.$$

Remark: The functions g_1 and g_2 could be allowed to depend on $\omega \in \Omega$ in a $\{\mathcal{G}_t\}$ -predictable way.

Assume that the $\{\mathcal{F}_t\}$ -compensator ν of Ψ is given by

$$\nu(dt \times \{y\}) = \mathbf{1}\{t < T'_\infty\} r(T^{t-}(\Psi^0), Y(t-), t, y) dt,$$

where

$$r : \Omega \times [0, \infty) \times \mathbf{Y} \times [0, \infty) \times \mathbf{Y} \rightarrow [0, \infty)$$

is a measurable function that is $\{\mathcal{G}_t\}$ -predictable in the first and fourth argument. Let

$$\bar{r}(s, z, u) := \sum_{y \in \mathbf{Y}} r(s, z, u, y).$$

and assume that

$$\int_s^t \bar{r}(s, z, u) du < \infty, \quad s \leq t,$$

$$\mathbb{E} \int_0^t \mathbf{1}\{u < T'_\infty\} \bar{r}(T^u(\Psi^0), Y(u), u) du < \infty, \quad t \geq 0,$$

Define

$$U(s, y, t) := \mathbf{1}\{s \leq t\} \exp \left[- \int_s^t \bar{r}(s, y, u) du \right].$$

Theorem: *We have*

$$\mathbb{P}(T^t(\Psi^0) \in ds, Y(t) = y | \mathcal{G}_t) = \frac{Q(t; y, ds)}{\sum_z Q(t; z, [0, \infty) \times \mathbf{Y})} \quad \mathbb{P} - a.s.,$$

where, for $n \geq 0$ and $T_n \leq t < T_{n+1}$

$$Q(t; y, ds) = C_n(y) \frac{U(0, y, t)}{U(0, y, T_n)} \delta_0(ds) + \mathbf{1}\{s \leq T_n\} p_n(s, y) \frac{U(s, y, t)}{U(s, y, T_n)} ds \\ + \mathbf{1}\{T_n < s \leq t\} p(s, y) U(s, y, t) ds.$$

The $C_n(y)$ are \mathcal{G}_{T_n} -measurable r.v.'s, $p_n : \Omega \times [0, \infty) \times \mathbf{Y} \rightarrow [0, \infty)$ is \mathcal{G}_{T_n} -measurable in the first argument, and $p : \Omega \times [0, \infty) \times \mathbf{Y} \rightarrow [0, \infty)$ is $\{\mathcal{G}_t\}$ -optional in the first two arguments. These functions are given by:

$$C_0(y) = P(Y_0 = y | \mathcal{G}_0),$$

$$C_{n+1}(y) = \sum_{z \in \mathbf{Y}} \mathbf{1}\{g_2(0, z, T_{n+1}, y) = X_{n+1}\} g_1(0, z, T_{n+1}, y)$$

$$r(0, z, T_{n+1}, y) C_n(z) \frac{U(0, z, T_{n+1})}{U(0, z, T_n)} \quad \text{on } \{T_{n+1} < \infty\},$$

$$p_0(s, y) = 0,$$

$$p_{n+1}(s, y) = \mathbf{1}\{s \leq T_n\} \sum_{z \in \mathbf{Y}} g_1(s, z, T_{n+1}, y) \mathbf{1}\{g_2(s, z, T_{n+1}, y) = X_{n+1}\}$$

$$r(s, z, T_{n+1}, y) p_n(s, z) \frac{U(s, z, T_{n+1})}{U(s, z, T_n)}$$

$$+ \mathbf{1}\{T_n < s \leq T_{n+1}\} \sum_{z \in \mathbf{Y}} g_1(s, z, T_{n+1}, y) \mathbf{1}\{g_2(s, z, T_{n+1}, y) = X_{n+1}\}$$

$$r(s, z, T_{n+1}, y) p(s, z) U(s, z, T_{n+1}), \quad T_n \leq t < T_{n+1}, n \geq 0.$$

Furthermore, we have for all $n \geq 0$ and all t with $T_n \leq t < T_{n+1}$,

$$\begin{aligned}
 p(t, y) &= \sum_{z \in \mathbf{Y}} (1 - g_1(0, z, t, y)) r(0, z, t, y) C_n(z) \frac{U(0, z, t)}{U(0, z, T_n)} \\
 &\quad + \sum_{z \in \mathbf{Y}} \int_0^{T_n} (1 - g_1(u, z, t, y)) r(u, z, t, y) p_n(u, z) \frac{U(u, z, t)}{U(u, z, T_n)} du \\
 &\quad + \sum_{z \in \mathbf{Y}} \int_{T_n}^t (1 - g_1(u, z, t, y)) r(u, z, t, y) p(u, z) U(u, z, t) du.
 \end{aligned}$$