

## Stochastic Processes Solutions to the 7th Problem sheet

### Solution to problem 1

Let  $Y_n = X_n$  for all  $n \in \mathbb{N}$ . This is equivalent to  $p(1) = P$ , i.e.  $p_{00}(1) = p_{11}(1) = \alpha$ .

In example 8.8 we derived that for a Markov chain  $Y = (Y_t)_{t \geq 0}$  with state space  $S = \{0, 1\}$  and generator

$$Q = \begin{pmatrix} -\lambda & \lambda \\ \mu & -\mu \end{pmatrix}, \quad \lambda, \mu > 0.$$

the transition probabilities are

$$p_{00}(t) = \frac{\lambda}{\lambda + \mu} e^{-(\lambda + \mu)t} + \frac{\mu}{\lambda + \mu},$$

$$p_{11}(t) = \frac{\mu}{\lambda + \mu} e^{-(\lambda + \mu)t} + \frac{\lambda}{\lambda + \mu}.$$

Since

$$p_{00}(1) = \alpha = p_{11}(1),$$

we have

$$\alpha = \frac{1}{2}(p_{00}(1) + p_{11}(1)) = \frac{1}{2} \left( 1 + e^{-(\lambda + \mu)} \right) > \frac{1}{2}.$$

Now let  $\alpha > \frac{1}{2}$ , then with  $\lambda := \mu := -\frac{1}{2} \log(2\alpha - 1)$  we have  $p_{00}(1) = p_{11}(1) = \alpha$  (and therefore the transition matrices are the same, i.e.  $Y_n = X_n$  for all  $n \in \mathbb{N}$ ).

Alternative proof: By

$$\frac{\lambda}{\lambda + \mu} e^{-(\lambda + \mu)} + \frac{\mu}{\lambda + \mu} = \alpha,$$

$$\frac{\mu}{\lambda + \mu} e^{-(\lambda + \mu)} + \frac{\lambda}{\lambda + \mu} = \alpha$$

we have

$$e^{-(\lambda + \mu)} + 1 = 2\alpha,$$

thus  $\lambda + \mu = -\log(2\alpha - 1)$ . Inserting, subtracting and applying  $2\alpha - 1 < 1$  proves the assertion and on top we have uniqueness of  $\lambda$  and  $\mu$ .

### Solution to problem 2

Define the matrices  $A$  and  $D$  by

$$A = \begin{pmatrix} \lambda & 1 \\ -\mu & 1 \end{pmatrix}, \quad D = \begin{pmatrix} -(\lambda + \mu) & 0 \\ 0 & 0 \end{pmatrix}$$

then  $Q = ADA^{-1}$  and

$$P(t) = \sum_{n=0}^{\infty} \frac{t^n}{n!} Q^n = A \left( \sum_{n=0}^{\infty} \frac{t^n}{n!} D^n \right) A^{-1} = A \begin{pmatrix} \exp(-t(\lambda + \mu)) & 0 \\ 0 & 1 \end{pmatrix} A^{-1}$$

$$= \frac{1}{\lambda + \mu} \begin{pmatrix} \lambda \exp(-t(\lambda + \mu)) + \mu & -\lambda \exp(-t(\lambda + \mu)) + \lambda \\ -\mu \exp(-t(\lambda + \mu)) + \mu & \mu \exp(-t(\lambda + \mu)) + \lambda \end{pmatrix}$$

### Solution to problem 3

The Kolmogorov forward differential equations are

$$p'_{ii}(t) = -\lambda p_{ii}(t), \quad p'_{ij}(t) = -\lambda p_{ij}(t) + \lambda p_{ij-1}(t)$$

The solution is

$$p_{ij}(t) = \begin{cases} 0 & , j < i \\ e^{-\lambda t} & , j = i \\ \lambda e^{-\lambda t} \int_0^t e^{\lambda s} p_{ij-1}(s) ds, & j > i \end{cases}$$

Now we prove inductively for  $j \geq i$

$$p_{ij}(t) = e^{-\lambda t} \frac{(\lambda t)^{j-i}}{(j-i)!}$$

i)  $j = i$  (✓)

ii)  $j \rightarrow j + 1$ :

$$p_{ij+1}(t) = \lambda e^{-\lambda t} \int_0^t e^{\lambda s} p_{ij}(s) ds = \lambda e^{-\lambda t} \int_0^t e^{\lambda s} e^{-\lambda s} \frac{(\lambda s)^{j-i}}{(j-i)!} ds = e^{-\lambda t} \frac{(\lambda t)^{j+1-i}}{(j+1-i)!}$$

Therefore  $P_0(X_t = i) = e^{-\lambda t} \frac{(\lambda t)^j}{j!}$ .

### Solution to problem 4

a)

$$\begin{aligned} P(X > t + s | X > t) &= \frac{P(X > t + s, X > t)}{P(X > t)} = \frac{P(X > t + s)}{P(X > t)} \\ &= \frac{e^{-\lambda_1(t+s)}}{e^{-\lambda_1 s}} = e^{-\lambda_1 t} = P(X > t) \end{aligned}$$

b)

$$P(X \wedge Y > x) = P(X > x, Y > x) = P(X > x)P(Y > x) = e^{-\lambda_1 x} e^{-\lambda_2 x} = e^{-(\lambda_1 + \lambda_2)x}$$

c)

$$\begin{aligned} P(X > Y) &= \int_0^\infty \int_v^\infty \lambda_1 e^{-\lambda_1 u} \lambda_2 e^{-\lambda_2 v} du dv = \int_0^\infty \lambda_2 e^{-\lambda_2 v} dv \int_v^\infty \lambda_1 e^{-\lambda_1 u} du \\ &= \int_0^\infty \lambda_2 e^{-\lambda_2 v} e^{-\lambda_1 v} dv = \frac{\lambda_2}{\lambda_1 + \lambda_2} \end{aligned}$$